

Euro Area Output Gaps and the Transmission of Common Shocks

Jointly modeling output gaps using a multicycle
Beveridge-Nelson decomposition.

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What are we doing?

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How are we doing it?

Propose multicycle trend-cycle decomposition using a prior on the signal-to-noise ratio.

Stabilization Policy and the Output Gap

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But how do we measure potential output and the output gap?

Beveridge Nelson Decomposition

- Assume we have a nonstationary time series y_t with a trend component τ_t that follows random walk with constant drift μ and a cyclical component c_t .

$$y_t = \tau_t + c_t \quad (1)$$

- Beveridge and Nelson (1981) define trend as

$$\tau_t = \lim_{j \rightarrow \infty} E_t[y_{t+j} - j \cdot \mu] \quad (2)$$

$$= y_t + \lim_{J \rightarrow \infty} \sum_{j=1}^J E[\Delta y_{t+j} - \mu]. \quad (3)$$

and cycle as

$$c_t = y_t - \tau_t \quad (4)$$

- We need to specify a **forecasting model** to estimate the long-horizon conditional expectation.

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- We separate choice of degree of shrinkage for full VAR system from targeting multiple cycles
- **Approach:** Specify a **prior on the signal-to-noise ratio** of the respective BN cycle of interest.
- **Intuition:** If included variables do not add information, system is shrunk towards univariate BN decomposition with a smooth trend

Data and Estimation

- Specify joint multicountry model¹ for four largest euro area economies²
- includes *real GDP*, *hours worked*, the *unemployment rate* and *inflation* for each economy from 2000Q1 to 2025Q2.
- Prior on signal-to-noise ratio is set to $\delta_0 = 10\%$.

► Shrinkage prior

¹16-variable BVAR(4)

²Germany, France, Italy, and Spain.

The Euro Area Output Gaps

Euro Area Output Gaps from Multicycle BN model

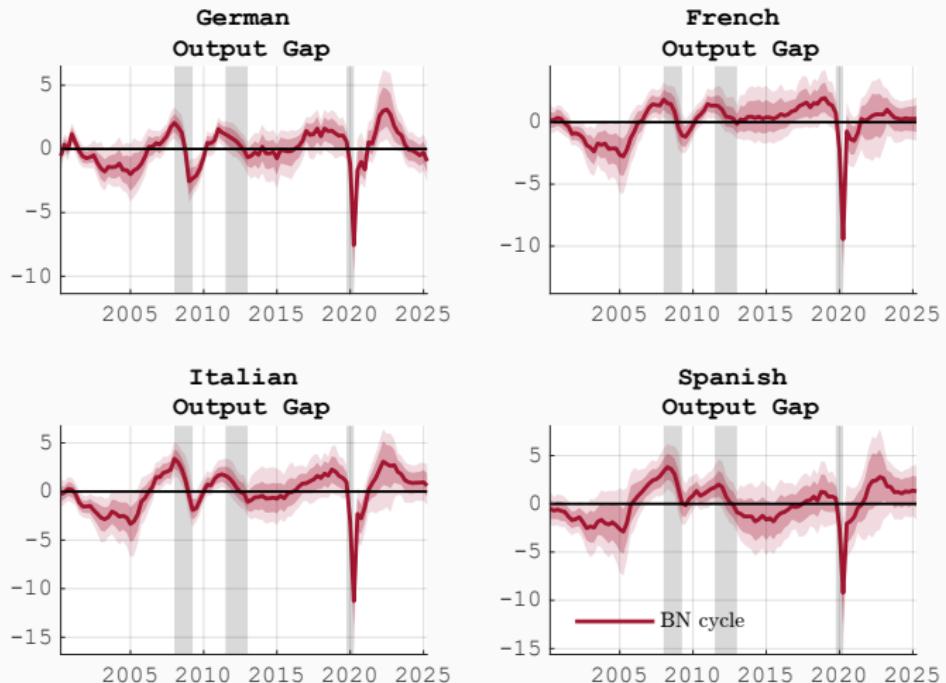


Figure 1: The euro-area output gaps. 90% credible sets given by dotted lines.

Informational decomposition of German Output Gap

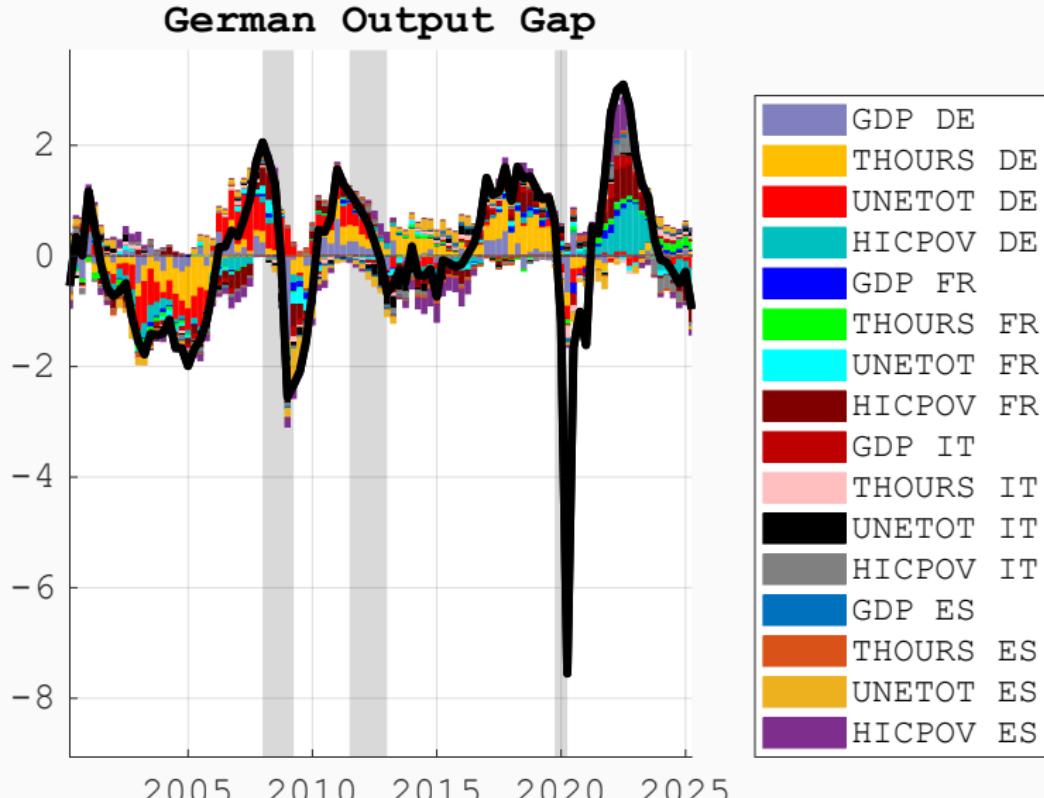


Figure 2: Information drivers of the German output gap

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- By writing each output gap $c_{i,t}$ as **function of structural shocks**, we can attribute dynamic causal effects

$$c_{ij,t} = - \sum_{l=0}^{t-1} \mathbf{s}_k \mathbf{F}^{l+1} (\mathbf{I} - \mathbf{F})^{-1} \mathbf{H} \mathbf{A} \mathbf{s}_j' \mathbf{s}_j \epsilon_{t-l}. \quad (5)$$

where $\mathbf{e}_t = \mathbf{A} \epsilon_t$ with $\mathbf{A} \mathbf{A}' = \Sigma$.

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- Obtain cross-correlations of euro area output gaps **conditional on a common shock**.

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- Obtain cross-correlations of euro area output gaps **conditional on a common shock**.
- Estimate **impulse responses** of euro area output gaps to common shocks.

Conclusion

Contribution

- Propose a **multicycle** extension to the BN decomposition.
- Present a joint model of euro area output gaps **fully accounting for cross-country dynamics/interlinkages**.

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- Propose a **multicycle** extension to the BN decomposition.
- Present a joint model of euro area output gaps **fully accounting for cross-country dynamics/interlinkages**.
- Analyzing the **transmission and spillovers of common shocks** across EA output gaps.

Thank you for your attention!

Data

The empirical application makes use of the large macroeconomic euro area dataset *EA_MD_QD* by Barigozzi and Lissone (2024). Table 1 gives an overview on the series used in the different specification and the respective data transformation. The *baseline* model only GDP growth and hours worked.

Table 1: Transformations. 1 - log level 2 - log differenced 3 - differenced.

Variable	Economy	Model	Transformation
Real GDP	EA9 economies	<i>baseline, robustness</i>	2
Empl	EA9 economies	<i>robustness</i>	2
Hours	EA9 economies	<i>baseline</i>	1
Unempl	EA9 economies	<i>baseline</i>	2
HICP	EA9 economies	<i>baseline</i>	2
Interest rate	Aggregate EA	<i>robustness</i>	0

Output gap measurement

Different approaches in the literature

- Estimation of potential output using production functions (Havik et al., 2014)
- Trend-cycle decompositions using statistical filter or state-space models (Hodrick and Prescott, 1997, Hamilton, 2018, Morley and Piger, 2012, Barigozzi and Luciani, 2023)

Our approach

- Potential output is output that should prevail in an economy **in absence of any cyclical shocks**
- The BN decomposition assumes the long-horizon conditional expectation of a time series only reflects trend

BN decomposition

Consider forecasting model

$$(\Delta \mathbf{X}_t - \mu) = \mathbf{F}(\Delta \mathbf{X}_{t-1} - \mu) + \mathbf{H}\mathbf{e}_t \quad (6)$$

Derive long-run forecast

$$E_t(\Delta \mathbf{X}_{t+1} - \mu) = \mathbf{F}(\Delta \mathbf{X}_t - \mu)$$

$$E_t(\Delta \mathbf{X}_{t+2} - \mu) = \mathbf{F}^2(\Delta \mathbf{X}_t - \mu)$$

$$\vdots \quad \vdots$$

$$E_t(\Delta \mathbf{X}_{t+j} - \mu) = \mathbf{F}^j(\Delta \mathbf{X}_t - \mu)$$

Given stationarity and eq. (6), we can write the cumulative sum at time t of expected future deviations of the vector process from its unconditional mean as

$$E_t \sum_{h=1}^{\infty} (\Delta \mathbf{X}_{t+h} - \mu) = \mathbf{F}(\mathbf{I} - \mathbf{F})^{-1}(\Delta \mathbf{X}_t - \mu). \quad (7)$$

Forecasting Model

- Set up medium scale Bayesian VAR as forecasting model

$$(\Delta \mathbf{X}_t - \mu) = \mathbf{F}(\Delta \mathbf{X}_{t-1} - \mu) + \mathbf{H}\mathbf{e}_t \quad (8)$$

- Hence, BN trend and cycle are given by (see Morley and Wong, 2020)

$$\tau_t = y_t + \mathbf{F}(\mathbf{I} - \mathbf{F})^{-1}(\Delta \mathbf{X}_t - \mu) \quad (9)$$

$$\mathbf{c}_t = -\mathbf{F}(\mathbf{I} - \mathbf{F})^{-1}(\Delta \mathbf{X}_t - \mu) \quad (10)$$

- Note each cycle $c_{i,t}$ fully decomposes into the forecast errors of all the K variables, i.e. $c_{i,t} = \sum_{j=1}^K c_{ij,t}$, where

$$c_{ij,t} = - \sum_{l=0}^{t-1} \mathbf{s}_k \mathbf{F}^{l+1} (\mathbf{I} - \mathbf{F})^{-1} \mathbf{H} \mathbf{s}_j' \mathbf{s}_j \mathbf{e}_{t-l}. \quad (11)$$

Shrinkage priors for VAR coefficients

Estimate VAR coefficient $\Phi_j, j = 1, \dots, p$ using **Minnesota-type shrinkage prior**. Specifically, we set

$$E[\phi_j^{*ik}] = 0 \quad (12)$$

$$Var[\phi_j^{*ik}] = \begin{cases} \frac{\lambda^2}{j^2} & , \text{ if } i=k \\ \frac{\lambda^2}{j^2} \frac{\sigma_i^2}{\sigma_k^2} & , \text{ otherwise.} \end{cases} \quad (13)$$

The Signal-to-Noise Ratio and the Output Gap

- Signal-to-noise ratio $\delta = \sigma_{\Delta_T}^2 / \sigma_{\varepsilon}^2$ is key concept in trend-cycle decompositions
- Lets assume an AR(p) as a competing forecasting model for Δy_t

$$\Delta y_t = \mu + \sum_{j=1}^p \phi_j (\Delta y_{t-j} - \mu) + \varepsilon_t, \quad \varepsilon_t \sim \mathcal{N}(0, \sigma_{\varepsilon}^2) \quad (14)$$

- There exists a direct mapping from the sum of autoregressive coefficients $\phi(1)$ to δ (see Kamber et al., 2018)

$$\delta = (1 - \phi(1))^{-2}$$

Informational and structural decomposition

We can derive the share of forecast error $c_{ij,t}$ of the j th variable in \mathbf{x}_t on the cycle $c_{i,t}$ while $\Delta y_{i,t}$ still has the k th position in \mathbf{x}_t

$$c_{ij,t} = - \sum_{l=0}^{t-1} \mathbf{s}_k \mathbf{F}^{l+1} (\mathbf{I} - \mathbf{F})^{-1} \mathbf{H} \mathbf{s}_j' \mathbf{s}_j \mathbf{e}_{t-l}. \quad (15)$$

Note that the cycle $c_{i,t}$ fully decomposes into the forecast errors of all the K variables contained in \mathbf{x}_t

$$c_{i,t} = \sum_{j=1}^K c_{ij,t} \quad (16)$$

To attribute dynamic causal effect in the interpretation of the cycle, we need to write $c_{i,t}$ as a function of orthogonal structural shocks

$$c_{ij,t} = - \sum_{l=0}^{t-1} \mathbf{s}_k \mathbf{F}^{l+1} (\mathbf{I} - \mathbf{F})^{-1} \mathbf{H} \mathbf{A} \mathbf{s}_j' \mathbf{s}_j \epsilon_{t-l}. \quad (17)$$

where $\mathbf{e}_t = \mathbf{A} \epsilon_t$ with $\mathbf{A} \mathbf{A}' = \Sigma$.

Cross-correlations of Output Gaps

- Recall the VAR model and the definition of the BN cycle

$$(\Delta \mathbf{X}_t - \mu) = \mathbf{F}(\Delta \mathbf{X}_{t-1} - \mu) + \mathbf{H}\epsilon_t$$

$$c_{i,t} = -\mathbf{s}_k \mathbf{F}(\mathbf{I} - \mathbf{F})^{-1}(\Delta \mathbf{X}_t - \mu)$$

- We know from Morley (2002) that $\mathbf{F}(\mathbf{I} - \mathbf{F})^{-1}(\Delta \mathbf{X}_t - \mu)$ contains the estimated BN cycles
- Berger et al. (2022) show that the variances of estimated BN cycles can be calculated as

$$\boldsymbol{\Psi} = \mathbf{F}(\mathbf{I} - \mathbf{F})^{-1} \boldsymbol{\Omega} [(\mathbf{I} - \mathbf{F})^{-1}]' \mathbf{F}'$$

where $\boldsymbol{\Omega}$ is the variance of $\Delta \mathbf{X}_t$ and $\text{vec}(\boldsymbol{\Omega}) = [\mathbf{I} - \mathbf{F} \otimes \mathbf{F}]^{-1} \text{vec}(\mathbf{Q})$, where

$$\mathbf{Q} = \begin{bmatrix} \boldsymbol{\Sigma} & \mathbf{0} & \dots \\ \mathbf{0} & \mathbf{0} & \ddots \\ \vdots & \ddots & \ddots \end{bmatrix}$$

- finally, normalize cross-covariance matrix $\boldsymbol{\Psi}$ into cross-correlation matrix ψ

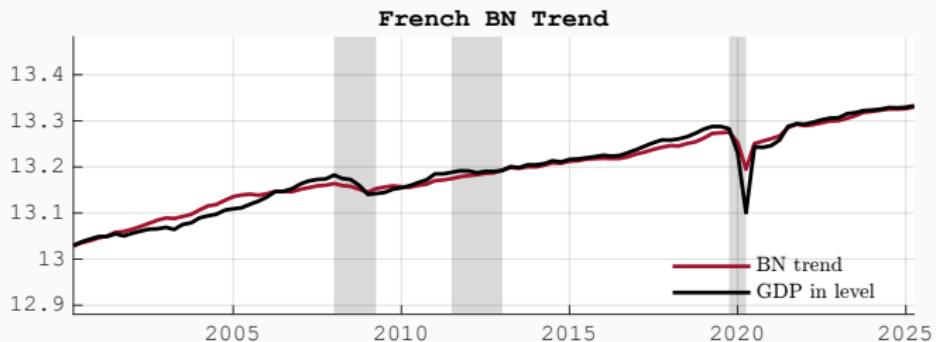
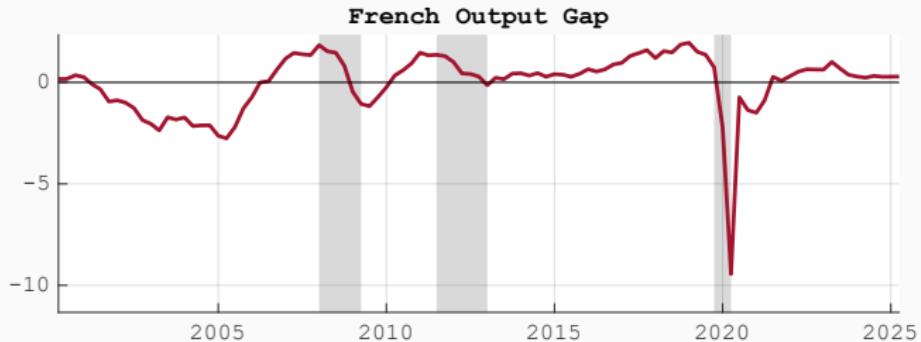
German Output Gap

Figure 3: German baseline results.



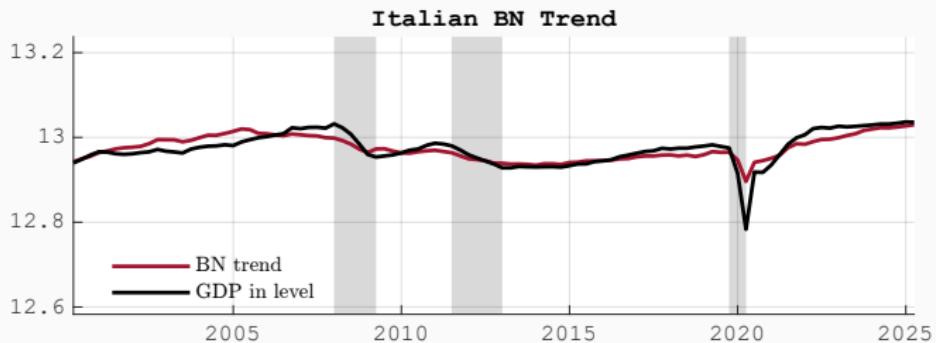
French Output Gap

Figure 4: French baseline results.



Italian Output Gap

Figure 5: Italian baseline results.



Spanish Output Gap

Figure 6: Spanish baseline results.

