

# Public Pressure and Heterogeneous Effects of Voluntary Pollution Abatement

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# Motivation

Voluntary pollution abatement programs (VPAs) are widely used complement for the mandatory regulations:

- Flexible: lack quantitative targets, sanctions and punishment.
- Do not require legislative action.
- No direct benefit.

# Research Questions

- How do VPAs create an incentive for firms to participate?
- What factors determine the effectiveness of VPAs in reducing pollution?

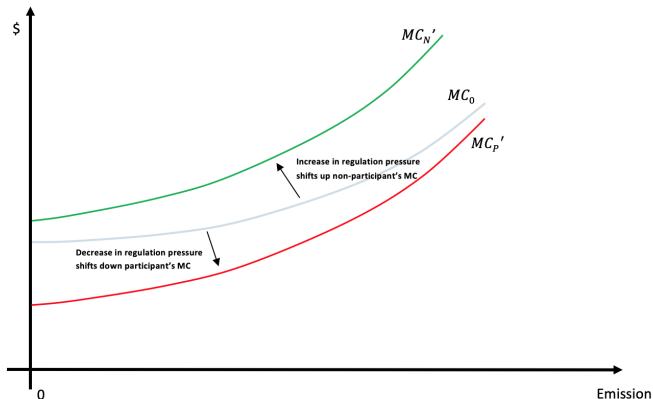
- VPAs are effective. (Khanna and Damon, 1999; Bi and Khanna, 2012; Innes and Sam, 2008).
- VPAs are not effective. (Welch et al., 2000; Gamper-Rabindran, 2006; Vidovic and Khanna, 2007; Brouhle et al., 2009; Carrion-Flores et al., 2013; Vidovic and Khanna, 2012; Vidovic et al., 2019)
- VPAs participants increase emissions compared to non-participants. (King and Lenox, 2000; Gamper-Rabindran and Finger, 2013)

# Three Factors Affecting Firm's Emission Behavior:

- Firm characteristics related to abatement cost
  - Example: Managerial skill, size, energy type, equipment, ownership...
- Emission cost from mandatory regulation pressure
  - Source: government intervention/regulation
  - Example: environmental taxes, expected inspections and violation penalties.
- Emission cost from public pressure
  - Source: Coasian bargaining in corporation-society conflicts ([Heal, 2005](#))
  - Example: “green” consumers boycott, loss in stock market, negative media coverage, local communities protests and lawsuits...

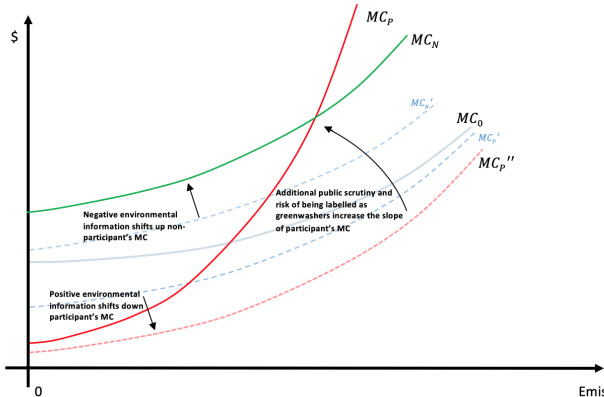
# Regulation Pressure

VPA shifts regulation resources from participants to non-participants ([Gamper-Rabindran, 2006](#); [Innes and Sam, 2008](#); [Li and Khanna, 2018](#)).



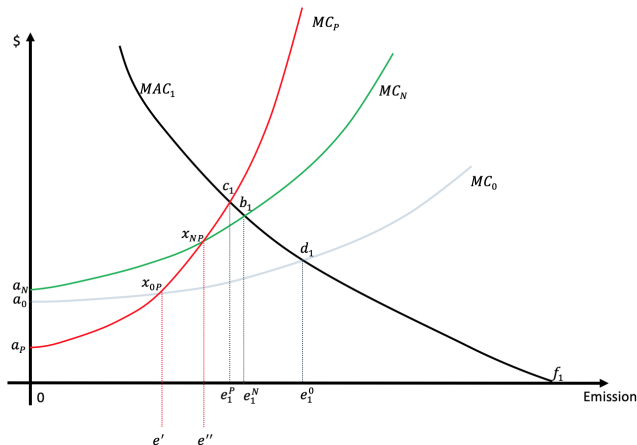
# Public Pressure

- VPA participation as positive news lowers public pressure, and vice versa (Aerts et al., 2008).
- VPA participation brings public scrutiny and the associated risk of being labeled greenwashers (Lyon and Maxwell, 2011; Kim and Lyon, 2011).



### 3 Types of Firms: High Pollution Firms

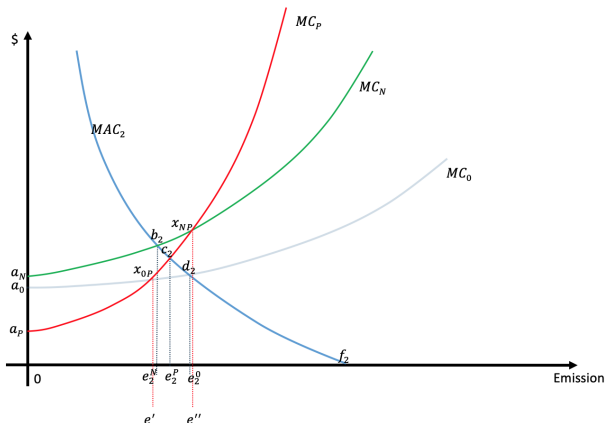
1.  $MAC$  intersects  $MC_P$  and  $MC_N$  on the right of  $x_{NP}$ ;
2. Always reduce emissions, participants ( $d_1 \rightarrow c_1$ ) reduce more than non-participants ( $d_1 \rightarrow b_1$ );
3. More polluting firms have smaller participation incentive ( $Area(a_N a_P x_{NP}) - Area(x_{NP} b_1 c_1)$ ;  $Area(x_{NP} b_1 c_1) \uparrow$ )





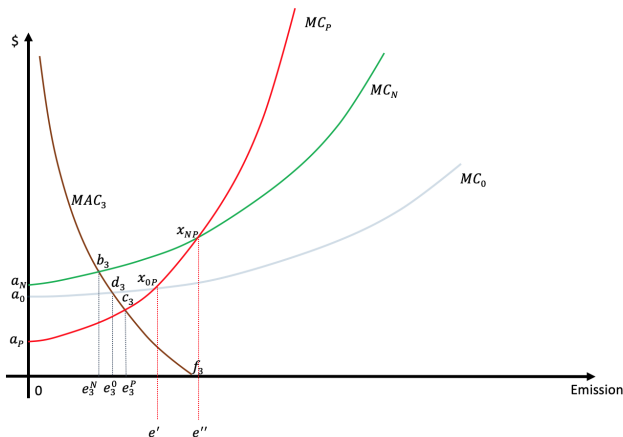
### 3 Types of Firms: Medium Pollution Firms

1.  $MAC$  intersects  $MC_P$  between  $x_{NP}$  and  $x_{OP}$ ;
2. Always reduce emissions, participants ( $d_2 \rightarrow c_2$ ) reduce less than non-participants ( $d_2 \rightarrow b_2$ );
3. More polluting firms have greater participation incentive ( $Area(a_n a_p c_2 b_2) \uparrow$ )



### 3 Types of Firms: Low Pollution Firms

1.  $MAC$  intercepts  $MC_P$  on the left of  $x_{OP}$ ;
2. Non-participants reduce emissions ( $d_3 \rightarrow c_3$ ), participants increase emissions ( $d_3 \rightarrow b_3$ );
3. More polluting firms have greater participation incentive ( $Area(a_n a_p c_3 b_3) \uparrow$ )



# Testable Hypothesis

**Hypothesis 1.** If the lagged emission level  $e_{t-1} \leq e''$ , then a marginal increase in  $e_{t-1}$  increases the probability of participation. If the lagged emission level  $e_{t-1} > e''$ , a marginal increase in  $e_{t-1}$  decreases the probability of participation.

**Hypothesis 2.** If the lagged emission level  $e_{t-1} \leq e''$ , then a participant reduces emissions by a smaller magnitude than non-participant (and is a free-rider). If the lagged emission level  $e_{t-1} > e''$ , then a participant reduces pollution by a larger amount than non-participant.

# Data

We use the EPA 33/50 program (1991-1995) data set from [Zhou et al. \(2020\)](#).

The goal of EPA 33/50 program was to reduce total emissions of 17 toxic chemicals by 33% by 1992 and by 50% by 1995, compared to the 1988 baseline.

There are 8,670 plants from 1988 to 1996.

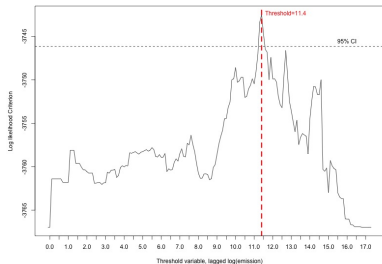
We consider two samples: the whole sample, and a 1991 sub-sample (with all non-participants and only first year (1991) participants).

# Empirical Model

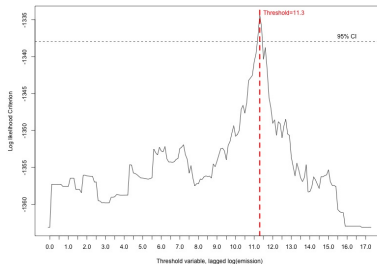
Two-stage regression to solve the endogeneity from self-selection in participation.

- First stage probit model estimates participation incentives
- Identify  $e''$  by grid search that maximizes likelihood of probit model (Hansen, 1999, 2000).
- Second stage dynamic panel model estimates the participation effect (instrumented by first stage estimates) using Arellano and Bond (1991) GMM.
- Allow heterogeneous participation effect on either side of the threshold  $e''$ .

# Empirical Results: First Stage



(a) MLE criterion of threshold parameter, 33/50 program participation incentive in 1991-1995



(b) MLE criterion of threshold parameter, 33/50 program participation incentive in 1991

Figure 1: Threshold estimation and 95% confidence interval

# Empirical Results: Second Stage

Table 1: Main analysis: 1991-1996 participation effect (benchmark results)

	<i>Dependent variable: log(33/50 emissions), first difference</i>					
	Full Sample	Full Sample		1991 Sub-Sample	1991 Sub-Sample	
		Before threshold	After threshold		Before threshold	After threshold
Participation Status	−0.556*** (0.053)	−0.183*** (0.062)	−0.956*** (0.104)	−0.287*** (0.072)	0.758*** (0.190)	−0.806*** (0.109)
Control Variables	Y		Y	Y		Y
Industry Specific Time Trends	Y		Y	Y		Y
State Specific Time Trends	Y		Y	Y		Y
Threshold log emission	-		11.4	-		11.3
Over-identification test (Hansen J Statistics P Value)	0.2794		0.3242	0.2105		0.2729
Weak identification test (Cragg-Donald Wald F Statistics)	69.611		70.083	49.133		47.385

Note: \*p<0.1; \*\*p<0.05; \*\*\*p<0.01

Covariates: lag log(33/50 emissions), HAP/TRI Ratio, LCV Score, county nonattainment status, county income per capita.

All covariates are in first differences.

IV's are first stage estimated participation probability and the third-year lag of log 33/50 emissions.

# Result Interpretation

$$\begin{aligned} e_{ijt}^P / e_{ij,t-1} &= e_{ijt}^N / e_{ij,t-1} \times \exp\left(\gamma_1 \mathbb{1}_{e_{ij,t-1} \leq e''} + \gamma_2 \mathbb{1}_{e_{ij,t-1} > e''}\right), \\ e_{ijt}^P &= e_{ijt}^N \times \exp\left(\gamma_1 \mathbb{1}_{e_{ij,t-1} \leq e''} + \gamma_2 \mathbb{1}_{e_{ij,t-1} > e''}\right). \end{aligned} \quad (1)$$

Participation effect estimated by the full sample:

- Below the threshold: -16.7%
- Above the threshold: -61.6%

Participation effect estimated by the 1991 Sub-Sample:

- Below the threshold: +113.4%
- Above the threshold: -55.3%



# Conclusion

- We develop a theoretical model illustrating a firm's incentives to participate in VPAs and outline the firm's emission decisions.
- The scrutiny of participating firms' environmental outcomes by the public is the key factor in determining whether VPAs participants free-ride the program or not.
- Public pressure can be effectively leveraged to complement traditional regulation and keep firms from free-riding in voluntary pollution abatement programs.

# Questions, Comments and Suggestions

Working paper available at:

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